Optimal Hedge Tracking Portfolios in a Limit Order Book

Simon Ellersgaard*

Department of Mathematical Sciences
University of Copenhagen
Universitetsparken 5
2100 Copenhagen, Denmark
s_ellersgaard@yahoo.com

Martin Tegnér
Mathematical Institute
University of Oxford, Woodstock Road
Oxford, OX2 6GG, England
†Oxford-Man Institute and Department of Engineering Science
University of Oxford, Parks Road
Oxford, OX1 3PJ, England
martin.tegner@eng.ox.ac.uk

Accepted 1 February 2018
Published 6 April 2018

Derivative hedging under transaction costs has attracted considerable attention over the past three decades. Yet comparatively little effort has been made towards integrating this problem in the context of trading through a limit order book. In this paper, we propose a simple model for a wealth-optimizing option seller, who hedges his position using a combination of limit and market orders, while facing certain constraints as to how far he can deviate from a targeted (Bachelierian) delta strategy. By translating the control problem into a three-dimensional Hamilton–Jacobi–Bellman quasi-variational inequality (HJB QVI) and solving numerically, we are able to deduce optimal limit order quotes alongside the regions surrounding the targeted delta surface in which the option seller must place limit orders vis-à-vis the more aggressive market orders. Our scheme is shown to be monotone, stable, and

*Corresponding author.
†Current affiliation.
consistent and thence, modulo a comparison principle, convergent in the viscosity sense.

Keywords: Delta hedging and limit order book; HJB QVI.

1. The Problem of Hedging with Transaction Costs

1.1. Introduction and literature survey

One of the core principles of prudent derivatives trading is the ability to place timely and accurate hedges of option portfolios. Reasons therefore extend beyond the curtailment of risk: by removing the delta-exposure of an option trade, a bet on implied versus realized volatility can be constructed, which in turn can be used to harvest volatility arbitrage (Ahmad and Wilmott, 2005; Ellersgaard et al., 2017). Yet, while hedging remains thoroughly understood in the axiomatic framework of Black–Scholes (1973), the shear empirical implausibility of this financial landscape in itself calls for an afterthought. Given that markets manifestly are not frictionless, any attempt at replicating a Black–Scholes hedge will theoretically result in infinite transaction costs (Zakamouline, 2006). It is precisely this uncomfortable fact which lies at the heart of much of the hedging literature produced over the past three decades. For example, Leland (1985), Hoggard et al. (1994) notably argue that for fixed-interval rebalancing of an underlying which incurs transaction costs in proportion to its price, the Black–Scholes equation takes on a modified volatility in order to match expected costs between derivative and replicating portfolio. While illuminating, this model nonetheless faces criticism from several directions. First, as shown by Kabanov and Safarian (1997), Leland’s conjecture of a vanishing hedging error is erroneous; secondly, as pointed out by Hodges et al. (Clewlow and Hodges, 1997; Hodges and Neuberger, 1989), Leland’s strategy clearly falls short of attaining any sense of optimality. In its stead, Hodges et al. ingeniously propose a utility indifference argument, in which the price of an option is determined as the amount of money required by a rational agent to make him indifferent to the prospect of being able to sell that option. While this approach uniquely fixes the price and optimal hedging strategy in a market with friction, it comes at the cost of diluting the principle of pure no-arbitrage pricing by the subjective utility preferences of the agent. Furthermore, as the Hamilton–Jacobi–Bellman (HJB) equation of the problem in itself is not analytically tractable, one must invariably resort

1 Specifically, if \( S_t \) is the mid-price of the underlying, and \( N \in \mathbb{Z} \) is the number of units acquired (sold) then \( \lambda |N| S_t \) is the spread paid for the pleasure, where \( \lambda \in \mathbb{R}_+ \).
to tedious numerical procedures as in Hodges and Neuberger (1989), or, more boldly, asymptotic solutions as championed by Whalley and Wilmott (1997). The latter duo, in particular, notes that for small proportional transaction costs, one effectively establishes a “no trade” region surrounding the original Black–Scholes delta, $\Delta_{BS}$, such that the option book is rebalanced just in case the current hedge ratio pierces the surrounding boundary. Although the associated analytic boundary captures some intuitively pleasing features such as enclosing on the delta when transaction costs go to zero, or when the agent’s risk aversion increases, it is important to appreciate that the Whalley–Wilmott result does not codify the full richness of the numerical solution. Indeed, as argued by Zakamouline (2004, 2006), the asymptotic approximation performs notably worse than its numerical counterpart for realistic model parameters.

Meanwhile, following Ho and Stoll’s (1981) trailblazing work from 1981 on optimal bid–ask quotes for wealth maximizing dealers, recent years have been marked by a veritable boom in research on optimal market making in limit order books (LOBs). Avellaneda and Stoikov (2008) have arguably played a pivotal role in facilitating this: by formalizing the control problem in the context of electronic market making — and justifying the arrival rate of market orders using principles of econo-physics — they effectively brought Ho and Stoll into the 21st century. Much momentum has subsequently been garnered, in part due to the contributions of Gueant et al. (2013), who transform the associated HJB equation into a system of linear ODEs, thereby enabling an actual solution, and Guilbaud and Pham (2011), who explore the inherent trade-off between being able to place “slow, but cost-effective” limit-orders, and “quick, but costly” market orders. Indeed, one cannot neglect Cartea and Jaimungal who have made a small industry out of the field with numerous noteworthy papers including (Cartea et al., 2013, 2014, 2016; Cartea and Jaimungal, 2014). Full justice to the rich literature on order-book prescriptivism cannot be done here; however, for readers interested in poignant surveys, alongside considerable original research, we refer to the excellent textbooks by Gueant (2016) and Cartea et al. (2015).

Surprisingly, comparatively little work has been done on trying to merge optimal hedging with the literature on dealings through a limit order book. To the best of our knowledge, only relatively few papers deal with this matter. Key citations here include Agliardi and Gencay’s (2014) discrete time investigation, in which an explicit solution is found for an option hedger who aims to minimise illiquidity costs and the hedging error, and Rogers and Singh’s (2010) determination of the optimal hedging policy in the face of
nonlinear transaction costs incurred as a function of the rate of change of the portfolio. Moreover, on the topic of pricing and hedging with temporary and permanent market impact, noteworthy papers include Li and Almgren (2016) and Gueant and Pu (2017). While superficially similar in nature, these papers employ fundamentally different philosophies in their interpretations of viable hedging: specifically, for the former, the “right” strategy is mean-reverting around the classical Black–Scholes delta (moving towards $\Delta_{BS}$ with a trading intensity proportional to the degree of mis-hedge and inversely proportional to illiquidity), while for the latter the hedging strategy in itself is changed to account for market micro-structure effects (ultimately resulting in smoother inventory paths, due to an in-built aversion towards round-trip trading).

In this paper, we aim to further bridge the gap between delta hedging and the order book literature. We do so from the perspective of a rational option seller, who derives linear utility from the financial wealth generated from his dealings, and quadratic utility penalization by deviating from some exogenously specified hedge target. Specifically, our concern here is the optimal policy the hedger should follow, insofar as the underlying asset can be traded using a combination of limit and market orders. The intuition here is that a hedger who is in close proximity to his delta target will be contended with placing slow and cost effective limit orders, while being significantly off-target will induce panic buying or selling through costly but instantaneous market orders. For us, the $\Delta$-strategy we are tracking is for all practical purposes completely generic, although we for illustrative reasons shall restrict our attention to a Bachelierian hedge ratio (Bachelier, 1900). Our framework thus finds its intellectual roots in the limit-market order duality developed by Guibaud and Pham (mentioned above), and Cartea and Jaimungal (2014), who consider an analogous set-up, albeit from the perspective of a trader who attempts to track the optimal execution algorithm developed by Almgren and Chriss (2000/2001). However, while the Almgren Chriss strategy in itself codifies market-microstructure effects by virtue of being sensitive to market impact, we stress that the same thing cannot be said of our targeted Bachelier hedge. In this sense, our philosophy sides with the work of Li and Almgren, again mentioned above.

More interesting variations of the targeted $\Delta$ are conceivable: for example, it would be opportune to explore tracking the Whalley Wilmott boundary, thereby facilitating entire regions of time-price-inventory space in which the hedger posts no orders at all. Thus, arguing that we are tracking the wrong delta fundamentally misses the point of what we are trying to accomplish. Moreover, while we submit that it ultimately is desirable to obtain a hedge
strategy endogenously within a model, we believe there is still room for a model like this. For example, it may be construed in the context of acting in compliance with externally enforced risk governance within a financial institution, or simply as stepping stones towards a better model, which is able to accommodate all of the opportunities that present themselves to option hedgers.

**Remark 1.** A plurality of sources provide highly readable accounts of the basic mechanisms and nomenclature of the limit order book, which prompts us to forgo a similar overview. Instead, we refer the reader to Gould *et al.* (2013), or Foucault *et al.* (2013).

1.2. Overview

The structure of the rest of this paper is as follows: in Sec. 2, we explicate the fundamental assumptions of the market and the portfolio manager and state the associated control problem. Section 3 is a survey of the exogenously specified hedge ratio: a careful analysis shows that the dynamics of the order book is weakly convergent towards the dynamics associated with the targeted delta. Section 4 exposes the Hamilton–Jacobi–Bellman quasi-variational inequality (HJB QVI) associated with the control problem, and suggests a dimensional reduction. Finding ourselves unable to extract an analytic solution thereto, Sec. 5 sets us up for a numerical scheme which is shown to have the desired properties of monotonicity, stability, and consistency as dictated by Barles (2012). Finally, Sec. 6 provides concrete numerical results, comparing the performance of our algorithm with a more naïve approach to $\Delta$-hedging. Section 7 concludes the paper.

**Remark 2.** We formulate our problem through the employ of Poisson point processes (see, e.g., Last and Brandt (1995) for a survey thereof). These constructs find their natural habitat through the inherent discreteness of the order book, and are instances of Hawkes processes, which have successfully been applied in the high-frequency domain of financial modeling in recent years, see, e.g., Bacry *et al.* (2015).

2. An Impulse-Control Approach to Hedging in the LOB

2.1. Market assumptions

We consider a financial market model carried by a filtered probability space $(\Omega, \mathcal{F}_T, \mathbb{F}, \mathbb{P})$, where the augmented filtration $\mathbb{F} = \{\mathcal{F}_t\}_{t \in [0,T]}$ is generated by
the random measures $J_1$, $J_2$ and $J_M$, which will be defined shortly. For simplicity, the market is assumed interest rate free and of a singular non-dividend paying risky asset (a stock) which is traded in a limit order book. All prices in the order book are assumed to take values from the discrete set $\mathcal{S} := \{\sigma_m n : n \in \mathbb{Z}\}$ where $\sigma_m > 0$ defines the smallest permissible price interval for orders within the book. In particular, we assume that the dynamics of the mid-price follows:

$$dS_t = \sigma_m \int_{\mathbb{R}} z(J_1(dt \times dz) - J_2(dt \times dz))$$

with initial value $s_0 = \sigma_m n$ for some positive integer $n$. The mid-price is driven by two independent Poisson random measures $J_1$ and $J_2$ on $[0, T] \times \mathbb{R}$ with common intensity measure $\mu_i(dt \times dz) = \gamma_i F_t(dz) dt$, for $i = 1, 2$. Here, $\gamma_i$ is a positive deterministic function which encodes the intensity of upwards (respectively downwards) jumps of the stock price at time $t$. Furthermore, $F_t$ is a discrete distribution function with support $\mathbb{Z}_+$ for each $t$ (with subscript + denoting the subset of positive values) that captures the conditional distribution of the number of tick points the price jumps, given that a jump occurs at time $t$.

The difference between the best bid, $S^b_t$, and the best ask, $S^a_t$, (the so-called bid–ask spread) is assumed time-dependent and of magnitude

$$S^a_t - S^b_t = 2\Upsilon_t,$$

where $\Upsilon_t$ is a deterministic function taking values in $\mathcal{S}_+$. Thus, investors who trade through market orders will minimally pay $S^a_t = S_t + \Upsilon_t$ per share if they wish to buy the stock, and maximally earn $S^b_t = S_t - \Upsilon_t$ per share if they wish to sell the stock, depending on the size of their orders vis-à-vis the number of shares available at the various price levels.

---

2 In practice, orders trading in the book are on a grid of integer multiples of the tick size, while the (non-tradable) “mid-price” is defined as the average of the best bid and best ask price, thus possibly taking values at half ticks where no orders are permitted. For simplicity, we treat all orders and prices to be taking values on the same set $\mathcal{S}$.

3 A couple of comments are worth attaching to (1): first, the dynamics is form-invariant upon rewriting it in terms of the compensated Poisson random measures $\tilde{J}_i(dt \times dz) = J_i(dt \times dz) - \mu_i(dt \times dz)$. Hence, $S_t$ is an $\mathcal{F}$-martingale. Secondly, there is a potential pathology nested in the dynamics in the sense that it admits negative price processes, i.e., $S_t \in \mathcal{S}$. Nonetheless, we shall assume that the probability of this occurring is sufficiently low to be ignored (an otherwise very reasonable assumption over short temporal horizons/for low intensity processes) or, at least, that all trading activity in the stock stops if the price reaches zero.

---

1850003-6
At the aggregate level, we assume that market orders arrive in the limit order book in a manner which can be modeled by an inhomogeneous Poisson process $M_t$, independent from $S_t$, with intensity rate

$$\lambda_t = \xi_t \exp\{\kappa_t \Upsilon_t\},$$

where $\xi_t$ and $\kappa_t$ are positive deterministic functions, and we denote the associated jump measure with $J_M$. Clearly, while market orders are guaranteed instantaneous execution, the liquidity taking fees they incur may be less than appealing. This is not to say that one optimally should opt for the limit order alternative: despite superior prices, there is no guarantee of execution at all.

Finally, we assume that the functions $\gamma_t$, $F_t$, $\Upsilon_t$ and $\lambda_t$ are sufficiently well behaved for the dynamics of the prices and order arrivals to be well defined; for instance, $\Upsilon_t$ being bounded and $\gamma_t \int \mathbb{R} z^2 F_t(dz) \leq K$ for a constant $K$ for all $t$, which makes $S_t$ a square-integrable martingale, and $\int_{[0,T]} (\gamma_t + \lambda_t) dt < \infty$ for the the jump rate intensities such that the inhomogeneous Poisson processes are well defined (requiring the intensity functions to be either bounded or continuous would suffice).

### 2.2. Portfolio manager assumptions

We consider the case of a portfolio manager who at time $t = 0$ sells off $N$ European call options of strike $K$ and maturity $T$ due for physical delivery. Modulo an exogenously specified tolerance for risk, his aim is to keep his portfolio “delta neutral” throughout, while simultaneously maximizing his terminal payoff at time $T$. Here, *delta neutrality* is to be understood as a pre-specified optimal inventory level $\Delta_t = \Delta(t, S_t)$ — the number of shares the portfolio manager ideally wishes to keep in his portfolio to hedge his short option position. If his inventory level at time $t$, denoted $Q_t$, is found to be in excess of $\Delta_t$, the portfolio manager is incentivized to offload some of his shares. Conversely, if $Q_t$ falls short of $\Delta_t$, he is incentivized to acquire shares. To this end, we assume that the portfolio manager places unit-sized limit orders and (if necessary) integer-sized market orders in the order book, but only one of the two at any given time. From this, one is prompted to ask the following questions, which will form the backbone of this paper:

**Question 1.** When should the portfolio manager trade in limit orders and when should he trade in market orders? In particular, regarding the former, which price quotes should he employ?
While market sell (buy) orders are assumed to take place at the best bid (ask), we assume the portfolio manager is at liberty to decide how deep in the limit order book he places his limit orders. Specifically, let $S^-_t = S_t + \delta^-_t$ be the price level at which the portfolio manager places his ask quote at time $t$, and let $S^+_t = S_t - \delta^+_t$ be the price level at which the portfolio manager places his bid quote at time $t$, where $\delta^\pm_t \in \mathcal{S}_+$ is known as the spread. For mnemonic purposes, we designate objects that give rise to a lower (higher) inventory by the superscript $- (+)$.

To capture the execution risk inherent to limit orders, we assume that the probability of a sell (buy) limit order being lifted, given that a market order arrives, is of the form $\exp\{-\kappa_t\delta^-_t\}$, $(\exp\{-\kappa_t\delta^+_t\})$. Thus, from the market assumptions, it follows from the thinning property of Poisson processes (of the market order arrivals $M_t$) that the portfolio manager’s successful limit sell and buy order executions are inhomogeneous Poisson processes $L^-_t$ and $L^+_t$, with intensities

$$
\lambda_t \exp\{-\kappa_t\delta^-_t\} \quad \text{and} \quad \lambda_t \exp\{-\kappa_t\delta^+_t\},
$$

respectively. The intuition here is clear: the further away from the mid-price the portfolio manager places his limit quotes, the less likely it is they will be executed. This dis-incentivizes the portfolio manager from posting extreme limit order quotes, even though he obviously would benefit considerably from their execution. Insofar as a limit order placed at time $t$ fails to be executed, we assume that the portfolio manager cancels it immediately, only to replace it with a (possibly) updated quote.

As the alternative to posting limit orders, let $0 < \tau^-_1 \leq \tau^+_2 \leq \cdots < T$ denote the stopping times at which the investor chooses to place market orders. We have that the total cash position $B_t \in \mathcal{S}$ (bank holding) and inventory level $Q_t \in \mathbb{Z}$ of the portfolio manager obey the jump formulae

$$
\begin{align*}
\frac{dB_t}{t} &= \mathbf{1}_{\{Q_t > \Delta_t\}}(S_t + \delta^-_t) dL^-_t - \mathbf{1}_{\{Q_t < \Delta_t\}}(S_t - \delta^+_t) dL^+_t, \\
\frac{dQ_t}{t} &= -\mathbf{1}_{\{Q_t > \Delta_t\}} dL^-_t + \mathbf{1}_{\{Q_t < \Delta_t\}} dL^+_t,
\end{align*}
$$

for $\tau^\pm_{i-1} < t < \tau^\pm_i$, i.e., in the intermediate time interval between market orders. Note here that the portfolio manager’s trading activity — as previously described in this section — effectively is constrained by the indicator functions: $\mathbf{1}_{\{Q_t > \Delta_t\}}$ codifies a sell initiative, potentially decreasing the inventory and increasing the cash position, while $\mathbf{1}_{\{Q_t < \Delta_t\}}$ codifies a buy initiative (potentially increasing inventory and decreasing cash). Further, a market order posted by the investor at time $\tau^\pm_t$ will bring the cash and
inventory to a new state

\[
B_{r_i}^+ = B_{r_i}^- + 1_{\{Q_{r_i}^- > \Delta_{r_i}^\pm\}}(S_{r_i}^+ - \Upsilon_{r_i}^\pm) - 1_{\{Q_{r_i}^- < \Delta_{r_i}^\pm\}}(S_{r_i}^- + \Upsilon_{r_i}^\pm),
\]

\[
Q_{r_i}^+ = Q_{r_i}^- - 1_{\{Q_{r_i}^- > \Delta_{r_i}^\pm\}} + 1_{\{Q_{r_i}^- < \Delta_{r_i}^\pm\}}.
\]  

A posting of two or more market orders at the time should be understood as a concatenation of the mapping \(\Gamma: (B_{t-}, Q_{t-}) \mapsto (B_t, Q_t)\) defined by (4), that is, \((B_{r_i}^\pm, Q_{r_i}^\pm) = \Gamma(B_{r_i}^\pm, Q_{r_i}^\pm)).\) Finally, as a result of the fact that \(0 \leq \Delta_t \leq \mathfrak{T}\), note that \(Q_t\) takes values in the finite set \(Q := \{0, \ldots, \mathfrak{T}\}\.\)

At maturity, we imagine that one of the following scenarios obtains: if the options are out of the money, the portfolio manager immediately liquidates his inventory using market orders. On the other hand, if the options are at/in the money, the option holders exercise their right to buy the stock for the strike price: insofar as there is a mismatch between the portfolio manager’s inventory and the \(\mathfrak{T}\) stocks due for delivery, he acquires the difference, again using market orders. Here, we note that given that expiry and delivery can be temporally separated events, it is conceivable that the portfolio manager would wish to acquire the difference using, say, the execution algorithm proposed by Cartea and Jaimungal (2014). Furthermore, had the options been cash settled, the portfolio manager would naturally be incentivized to offload, rather than acquire shares, in order to finance the claim made by his counter-parties (Gueant and Pu, 2017). These considerations present obvious ways in which our work may be extended.

Finally, suppose the portfolio manager derives linear utility from his level of financial wealth, but incurs a quadratic lifetime penalization from any deviation from the target hedge portfolio. The control problem to be solved can thus be stated as follows:

\[
V(t, b, s, q) = \sup_{\{\delta^-, \tau^+\} \in \mathcal{A}(t, b, s, q)} \mathbb{E}^{t, b, s, q}_t \left[ B_\theta + 1_{\{0 < S_\theta < K\}} Q_\theta (S_\theta - \Upsilon_\theta) \right. \\
+ 1_{\{S_\theta \geq K\}} [\mathfrak{N}K - (\mathfrak{N} - Q_\theta)(S_\theta + \Upsilon_\theta)] - \eta \int_t^\theta (Q_u - \Delta_u)^2 \, du \left. \right],
\]

where \(\mathbb{E}^{t, b, s, q}[]\) is the conditional expectation given \((B_t, S_t, Q_t) = (b, s, q) \in \mathcal{S} \times \mathcal{S}_+ \times \mathcal{Q})\), and the supremum runs over all admissible controls, \(\mathcal{A}(t, b, s, q)\), i.e., the set of all \(\mathbb{F}\)-predictable limit order spreads with values in \(\mathcal{S}_+\) sufficiently integrable for (3) to be well behaved, and all \(\mathbb{F}\)-stopping times bounded above by \(T\). Here, we have introduced the stopping time \(\theta := \min(\theta_S, T)\) where \(\theta_S = \inf\{u \geq t : S_u \not\in \mathcal{S}_+\}\) is the first time \(S_t\) reaches zero.
Hence, $\theta$ codifies that the investor may trade either until maturity, or until the game effectively is over, i.e., the event when the stock happens to reach zero and expires worthless. Finally, $\eta \in \mathbb{R}_+$ is a parameter which captures the portfolio manager’s “readiness” to depart from the desired hedge strategy $\Delta$ (clearly, the greater the $\eta$, the less prone the portfolio manager will be to depart from the prescribed strategy). In practical terms, $\eta$ may be seen as a function of compliance with external (regulatory) risk measures, as well as internal (company specific) risk management. Notice though, that analogous to more traditional risk aversion parameters, there is an inexorable nebulosity wrapped around this construct: whether real world risk preferences can be accurately mapped to this singular parameter, and if it can be done with a reasonable degree of empirical accuracy is arguably questionable.\(^4\)

3. The Question of the $\Delta$

The question of an optimal hedge strategy for jump processes in a market with friction is one of considerable complexity which we shall pass over in silence. As suggested above, our main concern here lies with the duality offered by the limit order book trading strategies; hence, the $\Delta$ we will consider is largely illustrative in nature. Specifically, in a slight variation of the work by Bachelier,\(^5\) we shall suppose that the portfolio manager aims to track a hedge strategy as though he were trading in a driftless arithmetic Brownian motion economy with time-dependent volatility, i.e., as though the market were frictionless with price dynamics

$$dS_t = \sigma_t dW_t,$$

where $W_t$ is a Wiener process and $\sigma : [0, T] \to \mathbb{R}_+$ is a deterministic function. Using a standard no-arbitrage argument, one may readily show that strike $K$ maturity $T$ call options should be priced according to

$$C^K_{t,T} = (S_t - K)\Phi(\delta_t) + \Sigma_t \phi(\delta_t),$$

where $\delta_t \equiv \delta(t, S_t) = \Sigma_t^{-1}(S_t - K)$, and $\Sigma_t \equiv (\int_t^T \sigma^2_u du)^{1/2}$, and we have introduced the usual functions: $\Phi(\cdot)$ as the standard normal cdf, and $\phi(\cdot)$ as the

\(^4\)Tentatively, one might try to Monte Carlo simulate portfolio returns under various $\eta$ specifications and study the associated portfolio return distributions.

\(^5\)Louis Bachelier is widely credited as the father of mathematical finance. In his doctoral dissertation *The Theory of Speculation* (Bachelier, 1900) Bachelier introduced Brownian motion in the modeling of stock prices. Specifically, he assumed arithmetic Brownian motion with constant parameters: $dS_t = \mu dt + \sigma dW_t$. 

1850003-10
standard normal pdf. Hence, from the net portfolio position $B_t + \Delta_t S_t - \mathcal{N} C_t$, to hedge a short position in $\mathcal{N}$ such call options, one should hold $\Delta_t = \mathcal{N} \partial \delta S C_t$ or, equivalently,

$$\Delta_t = \mathcal{N} \Phi(\delta_t),$$

(8)

units of the underlying asset, where we have used the standard identity $\phi'(x) = -x \phi(x)$. Viewed as a function of $(t, S_t)$, we note that $\Delta_t$ has the obvious properties that $\Delta \to \mathcal{N}$ when the options are deep in the money ($S_t \gg K$) and $\Delta \to 0$ when the options are deep out of the money ($S_t \ll K$). As $t \to T$, the transition between these two extremes becomes increasingly steep in its rise, ultimately converging towards the step function $\mathcal{N} \mathbb{1}\{S_T \geq K\}$ at expiry. Thus, the $\Delta$ of the money calls is exceedingly sensitive to fluctuations in the price process near expiry, potentially resulting in the acquisition or decumulation of a large amount of shares in a short span of time. An illustration of this $\Delta$ is provided in Fig. 1 for constant parameters.

Again, we emphasise that this choice largely is to get the ball rolling; the reader is encouraged to experiment with alternative specifications. However, besides the neglected issue of market friction, we also note that the hedge ratio (8) is not altogether groundless; specifically, we have the following noteworthy result:

**Proposition 1.** In the infinite intensity limit, the jump dynamics of the mid-price, (1), converges in distribution to the Bachelier dynamics, (6). In particular, for large $\gamma$ we have that

$$\sigma_t \approx \sigma_m \sqrt{2 E[Z^2]/\gamma},$$

(9)

where the over-line designates the function mean taken over the interval $[0, t]$.

**Proof.** We refer the reader to Appendix A for a proof of this result. \qed

**Remark 3.** The results presented here may be viewed as an abstract generalization of the textbook result that $(N_t - \gamma t)/\sqrt{\gamma}$ converges in distribution to $N(0, t)$ as $\gamma \to \infty$, where $N_t \sim \text{Pois}(\gamma t)$ cf., e.g., Eq. (2.81) p. 53 of Cont and Tankov (2004).

**Remark 4.** To accommodate the fact that only integer multiples of the underlying can be held (and thus, to avoid penalizing the portfolio manager unnecessarily), we shall in practice track the integer equivalent of the Bachelier delta, i.e., $\Delta_t = \lfloor \mathcal{N} \Phi(\delta_t) \rfloor$, where $\lfloor \cdot \rfloor$ designates the nearest integer function.
4. The Hamilton–Jacobi–Bellman Formulation

Following standard arguments in stochastic control theory, we postulate that the optimal value function (5) is the unique viscosity solution of the following HJB QVI:  \[ \max \left\{ \partial_t V + \gamma_t \int_{\mathbb{R}} [V(t, b, s + \sigma_m z, q) - 2V(t, b, s, q) + V(t, b, s - \sigma_m z, q)] F_t(dz) \right\} \]

\[ + \sup_{\delta^- \in S} \lambda_t e^{-\kappa \delta^-} [V(t, b + (s + \delta^-), s, q - 1) - V(t, b, s, q)] \mathbf{1}_{\{q > \Delta(t,s)\}} \]

\[ + \sup_{\delta^+ \in S} \lambda_t e^{-\kappa \delta^+} [V(t, b - (s - \delta^+), s, q + 1) - V(t, b, s, q)] \mathbf{1}_{\{q < \Delta(t,s)\}} \]

\[ - \eta (q - \Delta(t, s))^2; [V(t, b + (s - \Upsilon_t), s, q - 1) - V(t, b, s, q)] \mathbf{1}_{\{q > \Delta(t,s)\}} \]

\[ + [V(t, b - (s + \Upsilon_t), s, q + 1) - V(t, b, s, q)] \mathbf{1}_{\{q < \Delta(t,s)\}} \right\} = 0 \]

in \([0, T) \times S \times S_+ \times Q\), with the boundary conditions

\[ V(T, b, s, q) = b + \mathbf{1}_{\{0<s<K\}} q(s - \Upsilon_T) + \mathbf{1}_{\{s\geq K\}} \left[ \mathfrak{N} K - (\mathfrak{N} - q)(s + \Upsilon_T) \right], \]  \(10\)

\[ V(t, b, s, q) = b \quad \text{for} \quad s = 0. \]  \(11\)

A graphical representation of the terminal condition \(t = T\) is provided in Fig. 1 assuming \(B_T = 0\). Notice that the step function nature of the terminal hedge, \(\Delta_T\), clearly is reflected in the pay-off here: specifically, the value surface is constant along \(Q_T = 0\) for \(S_T < K\) and along \(Q_T = \mathfrak{N}\) for \(S_T \geq K\).

While searching for a closed form solution to (10) appears like an exercise in futility, it is immediately obvious that the problem at least offers a reduction in dimensionality. Specifically, since the money account \(B\) enters the

---

6 Verifying that the value function of an impulse control problem is the (unique) solution of a particular HJB QVI is in general a delicate task. One method of doing so (along with sufficient conditions for the existence and uniqueness of a viscosity solution for HJB QVIs) is given by Seydel (2009) for a general setting, where the state process is a jump-diffusion driven by Brownian motions and Poisson random measures. Based on these general results, Frey and Seydel (2010) consider a problem in a setting very similar to ours where the conditions of Seydel (2009) are certified to establish a solution for the value function. In particular, as noted in Frey and Seydel (2010), the continuity conditions expire for discrete state variables, which is the case of \((B_t, S_t, Q_t)\) in our setting. Thus, a thorough treatment of the necessary conditions in Seydel (2009) following in the lines of Frey and Seydel (2010) should eventually lead to guaranteeing the existence of a unique viscosity solution to (4), and that the solution is indeed the value function of our control problem (5). Since the level of technical detail of the conditions is beyond the scope of this paper, we refrain from giving a full proof here and refer the interested reader to Seydel (2009) and Frey and Seydel (2010).
Fig. 1. Assume $N = 100$ and $K = 50.5$, with constant values for $\sigma$ and $\Upsilon$. (a) The targeted Bachelierian hedge strategy $\Delta(t, S_t) = \Phi(\delta_t)$. Notice that the surface converges towards a step function at the expiry of the call options, (b) the terminal condition of the control problem disregarding the bank account. Observe that at the zeroth level of inventory we have the classic “hockey stick” pay-off structure associated with a short option position. At the other end, with a full level of inventory $Q = \mathfrak{R}$, the short option position is perfectly hedged which is reflected in the constant in-the-money pay-off.
terminal condition (10) linearly, and since the portfolio manager derives linear utility from his financial wealth, we have the ansatz

\[ V(t, b, s, q) = b + \theta(t, s, q), \]  

where \( \theta \) is a real valued function. Hence, the HJB QVI reduces to

\[ 0 = \max \left\{ \partial_t \theta(t, s, q) + \gamma_t \int_{\mathbb{R}} [\theta(t, s + \sigma_m z, q) - 2\theta(t, s, q) + \theta(t, s - \sigma_m z, q)] F_t(dz) \right. \]

\[ + \sup_{\delta_t^+ \in S} \lambda_t e^{-\kappa_t \delta_t^+} \left[ s + \delta_t^+ + \theta(t, s, q - 1) - \theta(t, s, q) \right] 1_{\{q > \Delta(t,s)\}} \]

\[ + \sup_{\delta_t^- \in S} \lambda_t e^{-\kappa_t \delta_t^-} \left[ -(s - \delta_t^-) + \theta(t, s, q + 1) - \theta(t, s, q) \right] 1_{\{q < \Delta(t,s)\}} \]

\[ - \eta(q - \Delta(t,s))^2 \left[ s - \Upsilon_t + \theta(t, s, q - 1) - \theta(t, s, q) \right] 1_{\{q > \Delta(t,s)\}} \]

\[ + \left[ -(s + \Upsilon_t) + \theta(t, s, q + 1) - \theta(t, s, q) \right] 1_{\{q < \Delta(t,s)\}} \right\}, \]  

subject to the boundary conditions

\[ \theta(T, s, q) = 1_{\{0 < s < K\}} q(s - \Upsilon_T) + 1_{\{s \geq K\}} [\mathfrak{M}K - (\mathfrak{M} - q)(s - \Upsilon_T)], \]

\[ \theta(t, s, q) = 0 \quad \text{for} \quad s = 0. \]  

Upon solving the first-order conditions

\[ \delta_t^{\pm} = \arg \max \lambda_t e^{-\kappa_t \delta_t^{\pm}} [s + \delta_t^{\pm} + \theta(t, s, q + 1) - \theta(t, b, s, q)], \]  

(15)

(where the sign \( \pm \) is chosen uniformly) and substituting these back in to (13), we obtain the following key result

**Proposition 2.** Assume \( \theta \) is a function which satisfies the following variational inequality

\[ 0 = \max \left\{ \partial_t \theta(t, s, q) + \gamma_t \int_{\mathbb{R}} [\theta(t, s + \sigma_m z, q) - 2\theta(t, s, q) + \theta(t, s - \sigma_m z, q)] F_t(dz) \right. \]

\[ + \sup_{\delta_t^+ \in S} \lambda_t e^{-\kappa_t \delta_t^+} 1_{\{q > \Delta(t,s)\}} + e^{-\kappa_t \delta_t^+} 1_{\{q < \Delta(t,s)\}} \]

\[ - \eta(q - \Delta(t,s))^2 \left[ s - \Upsilon_t + \theta(t, s, q - 1) - \theta(t, s, q) \right] 1_{\{q > \Delta(t,s)\}} \]

\[ + \left[ -(s + \Upsilon_t) + \theta(t, s, q + 1) - \theta(t, s, q) \right] 1_{\{q < \Delta(t,s)\}} \right\}, \]  

where the optimal limit order controls are given by

\[ \delta_t^{\pm} = \delta^{\pm}(t, s, q) = \kappa_t^{\pm} s + \theta(t, s, q) - \theta(t, s, q + 1), \]  

and the terminal condition is of the form (14), then the identity (12) yields a solution to the HJB QVI (10).
5. Towards a Numerical Solution

Qua the dimensionality reduction offered by (12), we have a three-variable variational inequality, which with limited computational expenditure can be solved using an explicit finite difference scheme. Here, we will elucidate how to set up the problem and demonstrate its convergence modulo a comparison principle (which would indeed be guaranteed subject to the conditions in Seydel (2009) discussed in the previous section). To this end, we invoke a standard argument along the lines of Barles (2012), Barles and Souganidis (1991) and Forsyth and Labahn (2007/2008), and show that our scheme is monotone, stable, and consistent. We remark that this is not the only way in which convergence of HJB numerical schemes may be demonstrated. For a probabilistic proof we refer the reader to Kushner and Dupuis (2001), whose seminal work on Markov chain approximations of the dynamic programming principle indubitably will resonate with some readers. For a viscosity approach to Markov chain convergence see Fleming and Soner (2005). For a highly readable account of finite difference methods in stochastic control theory we refer the reader to Tourin (2010).

Now, following standard procedure, to approximate (13) in the finite difference sense, we introduce the finite mesh

$$T^{\Delta t} \times S^\sigma_m \times Q = \{0, \Delta t, \ldots, n\Delta t, \ldots, N\Delta t = T\} \times \{s_{\min} = 0, \sigma_m, \ldots, i\sigma_m, \ldots, I\sigma_m = s_{\max}\} \times \{0, \ldots, j, \ldots, J = \mathcal{N}\},$$

in lieu of the full state space $[0, T] \times S_+ \times Q$. Here, $s_{\max}$ is an artificially imposed upper boundary which encapsulates the solution region of interest. Notice that the only discretization which takes place is along the time axis (the stock and inventory dimensions are manifestly already discrete in the original problem). Furthermore, to handle the integral in (13) suppose that the number of ticks that the mid-price can jump at any given time is bounded from above by a constant $K \in \mathbb{Z}_+$. Specifically, let $\int_{\mathbb{R}} F_s(dz) = \sum_{k=1}^K P_s\{Z = k\} = 1$ and $K \ll I$ such that the price process surely stays in the mesh unless $S_t < K\sigma_m$ or $S_t > (I - K)\sigma_m$. To block the possibility that the price process jumps out of the mesh near the boundaries, a suitable rescaling of the probability weights are performed in those regions. Finally, the explicit approximation to the HJB QVI may then be stated as

$$\vartheta_{ij}^n = \mathcal{G}_i[\vartheta^{n+1}] \equiv \max\{\mathcal{G}_i[\vartheta^{n+1}]; \mathcal{G}_0[\vartheta^{n+1}]\}, \quad (18)$$
for \( K \leq i \leq I - K \) and \( 1 \leq j \leq J - 1 \) where \( \vartheta_{i,j}^{n} \equiv \theta_{n\Delta t}(n\Delta t, i\sigma_{m}, j) \) is a finite difference approximation to the dimensionally reduced optimal value function \( \theta(t, s, q) \) at state \((n\Delta t, i\sigma_{m}, j)\), \( \vartheta_{i,j}^{n+1} \equiv (\vartheta_{i,j}^{n+1}, \vartheta_{i,j+1}^{n+1}, \vartheta_{i,j-1}^{n+1}, \{\vartheta_{i\pm s,j}^{n+1}\}_{s=1}^{K})^{T} \in \mathbb{R}^{2K+3} \), and we have defined the finite difference operators

\[
G_{\Delta t}[\vartheta_{i,j}^{n+1}] \equiv \vartheta_{i,j}^{n+1} + \Delta t \sum_{k=1}^{K} \mathbb{P}_{n+1}(Z = k)[\vartheta_{i+k,j}^{n+1} - 2\vartheta_{i,j}^{n+1} + \vartheta_{i-k,j}^{n+1}]
+ \sup_{\delta_{n+1}^{\pm} \in \mathcal{D}_{s}} \lambda_{n+1}e^{-\kappa_{n+1}\delta_{n+1}^{\pm}i\sigma_{m} + \delta_{n+1}^{\pm} + \vartheta_{i,j-1}^{n+1} - \vartheta_{i,j}^{n+1}}1_{(j > \Delta_{n+1})}
+ \sup_{\delta_{n+1}^{\pm} \in \mathcal{D}_{s}} \lambda_{n+1}e^{-\kappa_{n+1}\delta_{n+1}^{\pm}[-(i\sigma_{m} - \delta_{n+1}^{\pm}) + \vartheta_{i,j+1}^{n+1} - \vartheta_{i,j}^{n+1}]}1_{(j < \Delta_{n+1})}
- \eta(j - \Delta_{n+1})^{2},
\]

\[
G_{0}[\vartheta_{i,j}^{n+1}] \equiv [i\sigma_{m} - \gamma_{n+1} + \vartheta_{i,j-1}^{n+1}]1_{(j > \Delta_{n+1})} - [i\sigma_{m} + \gamma_{n+1} - \vartheta_{i,j+1}^{n+1}]1_{(j < \Delta_{n+1})}
\]

with analogous expressions near the boundaries. For completeness, note that the terminal condition of (18) takes on the following form:

\[
\vartheta_{i,j}^{N} = 1_{\{0 < i\sigma_{m} < K\}}j(i\sigma_{m} - \gamma_{N}) + 1_{\{i\sigma_{m} \geq K\}}\gamma_{K} - (\gamma_{n} - j)(i\sigma_{m} + \gamma_{N}), \quad (19)
\]

while the first-order conditions become

\[
\vartheta_{i,j}^{n+1} = \kappa_{n+1}^{-1}i\sigma_{m} + \vartheta_{i,j}^{n+1} - \vartheta_{i,j-1}^{n+1}. \]

Solving the control problem is thus algorithmically comparable to computing the no-arbitrage price of an American put option using finite difference methods: starting from the terminal condition and moving incrementally backwards in time, we must at each grid node decide whether a limit or a market order optimizes our expected utility.

**Proposition 3.** Assuming \( 2\gamma_{n}\Delta t \leq 1 \) for \( n = 1, 2, \ldots, N \), the numerical scheme (18), (19) is (i) monotone, (ii) stable and (iii) consistent. Thus, if the scheme is also satisfying a comparison principle, then it converges locally uniformly to the unique viscosity solution of (13), (14).

**Proof.** We refer the reader to Appendix B for a proof of this result. \( \square \)

6. Example

6.1. The compound Poisson model

With a converging numerical scheme at hand, we are in a position to compute the optimal limit order book hedge strategy called for in Question 1. For
simplicity, we will assume that all parameters are constant in time. In
particular, the dynamics of the mid-price is assumed to be the difference
between two compound Poisson processes:
\[
dS_t = \sigma_m (dY_1^t - dY_2^t)
\]
with \( S_0 = K \) and \( Y_i^t = \sum_{j=1}^{N_i^t} Z_j^i \), for \( i = 1, 2 \), where \{\( N_1^t, N_2^t \)\} \( t \in [0, T] \) are
independent Poisson processes of common intensity \( \gamma t \), while the jump
sizes \( \{Z_j^i\} \in \mathbb{Z}^+ \) are i.i.d. random variables assumed to follow the geometric
distribution\(^7\)
\[
\mathbb{P}\{Z_j^i = k\} = (1 - p)^{k-1} p.
\]
Here, \( k \in \mathbb{Z}^+ \), \( p \in (0, 1] \), and the first and second moments are of the form
\( \mathbb{E}[Z_j^i] = p^{-1} \) and \( \mathbb{E}[Z_j^i]^2 = (2 - p)p^{-2} \). Jumps of unit tick size thus dominate
the price process fluctuations, while larger jumps occur at a power
decaying rate.\(^8\) Finally, as a direct consequence of (8) alongside the approxima-
tion (9), the targeted Bachelierian hedge strategy must be of the
following form:
\[
\Delta_t = \left[ \mathfrak{N}(\Phi - \frac{S_t - K}{\sigma_m \sqrt{2\gamma p^{-2}(2 - p)(T - t)}}) \right]. \quad (20)
\]

6.2. Simulation

Definition 1. We designate the joint process \( \{S_t, Q_t\}_{t \in [0, T]} \) (i.e., the price
path traced out by the underlying asset over \( [0, T] \), together with the
associated inventory level held by the portfolio manager) the stock-
inventory path.

Remark 5. Clearly, while the stock path is exogenously determined by
market forces, the portfolio manager is at liberty to choose his associated
inventory level.

\(^7\)Readers familiar with Lévy processes will notice that this essentially is a discrete symmetric
version of the Kou model:
\[
\nu(dx) = [p \lambda_+ e^{-\lambda_+ x} \mathbf{1}_{\{x > 0\}} + (1 - p) \lambda_- e^{-\lambda_- x} \mathbf{1}_{\{x < 0\}}] dx,
\]
where \( \lambda_+, \lambda_- > 0 \) and \( p \in [0, 1] \).

\(^8\)In practical applications, we cut off the mass function at a level \( k \) such that the probability of
a jump of size \( k \) occurring is much less than one over the total number of incremental time
steps \( N \).
Consider the parametric specifications listed in Table 1. Upon solving the HJB QVI numerically, we find the optimal trade regions exhibited cross-sectionally in Fig. 2 for various times to maturity. Jointly, these temporal hypersurfaces form the surface plots exhibited in Fig. 3. The interpretation is as follows: the portfolio manager aims to keep his portfolio approximately delta-neutral understood in the sense of keeping his stock-inventory path on a

Fig. 2. (Color online) The optimal limit/market order regions for the portfolio manager at different times to maturity. The white line which cuts each figure across diagonally is the targeted inventory level $\Delta_t$ for different values of the underlying stock. The black region surrounding it is the stock-inventory level for which the portfolio manager is contented with placing limit orders in an attempt to attain delta neutrality. Finally, the light gray region (top left) and the dark gray region (bottom right) correspond to market orders, i.e., are market order regions in which the portfolio manager deems it necessary to trade immediately to bounce back into the limit order region. For the corresponding surface plots see Fig. 3.

Table 1. Parameter specifications used in simulation experiment. Some arbitrariness surrounds these numbers in the sense that no empirical calibration has been performed — nonetheless, they serve quite nicely for illustrative purposes.

<table>
<thead>
<tr>
<th>Quantity</th>
<th>Interpretation</th>
<th>Magnitude</th>
<th>Quantity</th>
<th>Interpretation</th>
<th>Magnitude</th>
</tr>
</thead>
<tbody>
<tr>
<td>$K$</td>
<td>Strike</td>
<td>50.5</td>
<td>$\eta$</td>
<td>Risk aversion</td>
<td>1</td>
</tr>
<tr>
<td>$\sigma_m$</td>
<td>Tick-size</td>
<td>0.5</td>
<td>$\kappa$</td>
<td>Limit decay</td>
<td>0.3</td>
</tr>
<tr>
<td>$T$</td>
<td>TTM</td>
<td>1</td>
<td>$\lambda$</td>
<td>Limit intensity</td>
<td>100</td>
</tr>
<tr>
<td>$N$</td>
<td>No. of calls</td>
<td>100</td>
<td>$\gamma$</td>
<td>Market intensity</td>
<td>200</td>
</tr>
<tr>
<td>$\Upsilon$</td>
<td>Half-spread</td>
<td>1</td>
<td>$p$</td>
<td>Jump size prob.</td>
<td>0.9</td>
</tr>
</tbody>
</table>
par with the hedge surface $\Delta_t$ (whatever the realized value of the underlying, $S_t$, may be). However, he is contended with merely posting limit sell or buy orders (depending on whether he is above or below $\Delta_t$) in the immediately adjacent region marked by black in Fig. 2, thus allowing for minor deviations from the targeted hedge ratio. Should his stock-inventory path pierce either of the gray surface, his “risk aversion” kicks in, and he performs an immediate inventory alteration through unfavorable market orders (to the point where he re-enters the limit order region). Notice that as the market sell and the
market buy surfaces both converge towards the strike $K$ at maturity, it becomes increasingly difficult to avoid placing market orders with time if $S_t \approx K$.

The implications of this are forcefully demonstrated in Figs. 4 and 5 for the realization of a single stock path. Here, we investigate the performance of the portfolio manager’s combined limit-market order hedge strategy vis-à-vis the naïve hedge strategy which at all times utilizes market orders to maintain delta neutrality. Unsurprisingly, since the former strategy permits all stock-inventory paths sandwiched between the optimal market surfaces, while the latter only allows those stock-inventory paths which are embedded in the optimal hedge surface, $\{Q_t\}_{t\in[0,T]}$ and $\{B_t\}_{t\in[0,T]}$ are considerably less erratic in case of the former. Indeed, the lower frequency of trading, combined with the favorable prices associated with limit orders, yield a much more favorable cash position at the expiry of the options. Specifically, for the simulation at hand, the number of (not necessarily unit sized) market orders placed with the naïve strategy amounts to 216, while the number of market-limit orders placed with the control strategy come in at 37 and 30, with an average spread for the limit orders of 2.79. As for the mean quadratic variation of the inventory path, $\overline{(\Delta Q)^2} \equiv \frac{1}{N} \sum_{i=1}^{N} (\Delta Q_i)^2$, the naïve strategy amounts to 1.768, while the control strategy comes in at 0.205. The analogous quantity, computed for the bank path yields 4424 versus 472.

To develop an understanding of the distributive properties of the hedge strategies above, we repeat the experiment $10^4$ times. The resulting histograms are exhibited in Fig. 6, while the estimators are given in Table 2. As expected, the distribution of the terminal stock value (Fig. 6(a)) falls symmetrically around the strike in a normally distributed manner, with the best fit being obtained around the tails. This highlights that the employ of the Bachelier dynamics in lieu of the jump processes is not altogether innocuous. As for the rate of return of the bank account, $B_T/B_0$, we refer to Fig. 6(b). While it is clear that the combined limit market order strategy is superior to the naïve market order strategy, the “camel hump” nature of the

---

9 In simulating a Poisson path, it is helpful to recall that the time spent between consecutive jumps, $\tau$, is a random variable which follows the exponential distribution $P\{\tau \leq t\} = 1 - \exp(-\gamma t)$, see Privault (2017) and Shreve (2010).

10 Since we assume that the portfolio manager only buys and sells integer quantities of the underlying, it is clear that the embedding should be read as “correct to the nearest integer”.

11 This is hardly surprising: e.g., it is a well-known fact that the Skellam distribution (the law of the difference between two Poisson random variables, $N_1 - N_2$) tends to the normal distribution if (i) $N_1$ and $N_2$ have common intensity and, importantly, (ii) $N_1 - N_2 = k$ is large. See Abramowitz and Stegun (1965).
Fig. 4. (Color online) Instantiation of a stock-inventory path. (a) The evolution of the mid-price, seeded at \( S_0 = K \), (b) the evolution of the portfolio manager’s inventory when he hedges using market orders only (gray) or a combination of market and limit orders (black), (c) the evolution of the portfolio manager’s bank account, again using market orders only (gray) or a combination of market and limit orders (black). Notice the latter outperforms the former. The bank wealth is (arbitrarily) seeded at \( B_0 = 10^4 \), (d) the magnitude of the orders placed by the portfolio manager when he hedges with market orders only, (e) the magnitude of the orders placed by the portfolio manager when he hedges with a combination of market (gray) and limit (black) orders, (f) the size of the spread posted by the portfolio manager in those cases where his limit orders are successful. Black stems are sells, while gray stems are buys.
The histogram might *prima facie* strike one as puzzling. This mystery is quickly deflated, however, upon realizing that the terminal value of the targeted hedge strategy — the step function $\mathcal{R}\mathcal{I}\{S_T \geq K\}$ — invariably will drive the portfolio manager to either liquidate or acquire $\sim 50$ shares relative to his initial inventory. The asymmetry between the negative and positive return humps is likely to be symptomatic of the asymmetric boundary condition (i.e., the operations executed given that the calls are in or out of the money). Finally, Fig. 6(c) exhibits the return of the hedge portfolio, $\Pi_T/\Pi_0$.

Fig. 5. (Color online) (a) A three-dimensional view of the stock-inventory path described in Fig. 4 when the portfolio manager uses a combination of limit and market orders. Importantly, the path does not remain embedded in the gridded hedge surface, but is free to bounce back and forth between the encompassing market surfaces discretely highlighted in transparent gray, (b) the graph viewed from $Q = 9t$. Notice the symmetric nature of the market surfaces viewed from this angle, (c) the graph viewed from $t = 0$. Here, it is quite apparent that the path moves within the pocket space defined by the surrounding market surfaces.
where

$$\Pi_t = B_t + \Delta_t S_t - \mathfrak{M} \cdot C_t.$$  

Blatantly, $\Pi_{t<T}$ is a model-dependent quantity which prompts us to suggest the at-the-money Bachelier price of the call option,

$$C_{t,\text{atm}}^{K,T} = \sigma_m \sqrt{\frac{\gamma(2-p)(T-t)}{\pi p^2}},$$
in accordance with Eq. (7). Again, while this is dynamically consistent with the choice of the $\Delta_t$, (20), it is also in flagrant disregard for the jump nature of the underlying and the friction prevailing on the market. Hence, the Bachelier specification $\Pi_B^0$ should not be seen as an absolute standard. Nevertheless, it is intriguing to note that it on average yields a positive return for the portfolio manager who uses a combination of limit and market orders to hedge his position, while the naïve hedge engenders a mean negative return. Indeed, a Welchian $t$-test shows that we can comfortably dismiss the null hypothesis of equal means at the 99% level. For more general values of $\Pi_0$, we simply observe that the mean of the control strategy return is higher than that of the naïve strategy, while the opposite is the case for the standard deviation.

7. Conclusion

The emergence of certain stylized properties of the otherwise hyper-complex limit order book beckons us to attempt mathematical modeling thereof. In this paper, we proposed a simple jump model of the mid-price, alongside a probability of having a limit order met which is an exponentially decaying function of the distance to the mid-price. It was shown that the jump model converges weakly to Bachelier’s arithmetic Brownian motion for large values of the intensity. Furthermore, under the assumption that a portfolio manager wishing to $\Delta$-hedge his short call position derives linear utility from his terminal wealth, but incurs quadratic lifetime penalization from deviating from a targeted hedge strategy, we derived numerical values for his optimal limit order quotes and stopping times at which he should switch to pure market orders. Following the standard literature on convergence in the viscosity sense, we proved that our scheme is monotone, stable, and consistent. We

<table>
<thead>
<tr>
<th></th>
<th>$B_T/B_0$ LIMs</th>
<th>$B_T/B_0$ MOs</th>
<th>$\Pi_T/\Pi_B^0$ LIMs</th>
<th>$\Pi_T/\Pi_B^0$ MOs</th>
</tr>
</thead>
<tbody>
<tr>
<td>Mean</td>
<td>0.9554</td>
<td>0.8868</td>
<td>1.0048</td>
<td>0.9445</td>
</tr>
<tr>
<td>Standard dev.</td>
<td>0.2345</td>
<td>0.2521</td>
<td>0.0077</td>
<td>0.0199</td>
</tr>
<tr>
<td>1% Quantile</td>
<td>0.6856</td>
<td>0.5828</td>
<td>0.9848</td>
<td>0.9044</td>
</tr>
<tr>
<td>25% Quantile</td>
<td>0.7183</td>
<td>0.6256</td>
<td>1.0000</td>
<td>0.9292</td>
</tr>
<tr>
<td>50% Quantile</td>
<td>0.9549</td>
<td>1.0574</td>
<td>1.0052</td>
<td>0.9438</td>
</tr>
<tr>
<td>75% Quantile</td>
<td>1.1926</td>
<td>1.1327</td>
<td>1.0102</td>
<td>0.9598</td>
</tr>
<tr>
<td>99% Quantile</td>
<td>1.2119</td>
<td>1.1763</td>
<td>1.0210</td>
<td>0.9843</td>
</tr>
</tbody>
</table>
also exemplified the utility of our algorithm by comparing it to a naïve strategy which deploys nothing but market orders: a clear augmentation of the mean return and a clear reduction in the associated variance were established here. Finally, given the generic nature of our model, we reiterate that it is adaptable to more sophisticated expressions for the targeted hedge ratio in the limit order book.

Acknowledgments

The authors are grateful to Mark H. A. Davis (Imperial College London), Nicole Branger (University of Münster), Charles-Albert Lehalle (Capital Fund Management), and two anonymous referees who made detailed and helpful comments on various iterations of this paper.

Ellersgaard (corresponding author) gratefully acknowledges support from the Danish Strategic Research Council, Program Committee for Strategic Growth Technologies, via the research center HIPERFIT: Functional High Performance Computing for Financial Information Technology (hiperfit.dk) under contract number 10-092299. Tegnér gratefully acknowledges financial support from the Wallenberg Foundations. The authors declare no conflict of interest.

Appendix A. Proof of Proposition 1

To establish the convergence of the jump dynamics to the Bachelier dynamics for large jump intensities, we invoke the following lemma:

Lemma A.1. Let \( L_t^i \equiv L_t^1 - L_t^2 \), where

\[
L_t^i \equiv L_t^i((0, t] \times \mathbb{R}) = \int_0^t \int_\mathbb{R} z J_i(dt \times dz)
\]

for \( i = 1, 2 \). Furthermore, let us introduce the function \( \zeta : [0, T] \mapsto \mathbb{R}_+ \) by

\[
\zeta_t \equiv \left( 2 \int_0^t \mathbb{E}[Z_s^2] \gamma_s ds \right)^{-1/2},
\]

(A.1)

where \( \mathbb{E}[Z_s^2] = \int_\mathbb{R} z^2 F_s(dz) \) and assume that \( \sup_{s \in [0, t]} \int_\mathbb{R} z^4 F_s(dz) < \infty \). Then, we have the following convergence in distribution:

\[
\zeta_t L_t \overset{d}{\to} N(0, 1) \quad \text{as} \quad \int_0^t \gamma_s ds \to \infty,
\]

(A.2)

where \( N(0, 1) \) is the standard normal distribution.
Proof. By Levy’s Continuity Theorem, it suffices to show that the characteristic function \( \zeta_t L_t \) converges pointwise to the characteristic function of \( N(0,1) \). To this end, suppose we partition the time interval into \( n+1 \) segments: \( T_n = \{ t_0, t_1 = t_0 + \Delta t_0, t_2 = t_1 + \Delta t_1, \ldots, t_{n+1} = t = t_n + \Delta t_n \} \), where \( \max_i \Delta t_i \to 0 \) as \( n \to \infty \). Similarly, let \( A_m = \{ \Delta A_1, \Delta A_2, \ldots, \Delta A_m \} \cup \{ \Delta A_j \} \) be a partition of the mark space over disjoint subsets. From the elementary properties of Poisson random measures it follows that \( J^q_i \equiv J((t, t_s + \Delta t_s) \times \Delta A_q) \) and \( J^{ur}_i \equiv J((t, t_u + \Delta t_u) \times \Delta A_r) \) are independent provided that \( q \neq r \). Furthermore,

\[
P\{J^q_i = k\} = e^{-\mu^q_i} \frac{\left(\mu^q_i\right)^k}{k!},
\]

where we have defined the discrete intensity measure \( \mu^q_i = \mu_i((t, t_s + \Delta t_s), \Delta A_q) = \gamma_s F_i(\Delta A_q) \Delta t_s \). Thus, the characteristic function of \( \zeta_t L_t \), i.e.,

\[
\varphi_{\zeta_t L_t}(a) \equiv \mathbb{E}[\exp\{ia \zeta_t L_t\}]
\]

where \( \varphi_{\zeta_t L_t}(a) \) is continuous on \( 0 \), and if \( \varphi \) is continuous on \( 0 \), and \( a \in \mathbb{R} \), can be decomposed as \( \varphi_{\zeta_t L_t}(a) \varphi_{L_t^a}(-a) \) (by the independence of \( L^1 \) and \( L^2 \)) whence

\[
\varphi_{\zeta_t L}(a) = \mathbb{E}\left[\exp\left\{ia \zeta_t \int_0^t \int_A zJ_1(ds \times dz)\right\}\right] \mathbb{E}\left[\exp\left\{-ia \zeta_t \int_0^t \int_A zJ_2(ds \times dz)\right\}\right]
\]

\[
= \lim_{m,n \to \infty} \mathbb{E}\left[\exp\left\{ia \zeta_t \sum_{s=0}^n \sum_{q=1}^m z_p J^q_1\right\}\right] \mathbb{E}\left[\exp\left\{-ia \zeta_t \sum_{u=0}^n \sum_{r=1}^m z_r J^{ur}_2\right\}\right]
\]

\[
= \lim_{m,n \to \infty} \prod_{s,u=0}^n \prod_{q,r=1}^m \mathbb{E}[\exp\{ia \zeta_{n+1} z_q J^q_1\}] \mathbb{E}[\exp\{-ia \zeta_{n+1} z_r J^{ur}_2\}]
\]

\[
= \lim_{m,n \to \infty} \prod_{s,u=0}^n \prod_{q,r=1}^m \sum_{k_1, k_2=0}^{\infty} \mathbb{P}\{J^q_1 = k_1\} \mathbb{E}\{ia \zeta_{n+1} z_q J^q_1\} | J^q_1 = k_1\}
\]

\[
\cdot \sum_{k_2=0}^{\infty} \mathbb{P}\{J^{ur}_2 = k_2\} \mathbb{E}\{\exp\{-ia \zeta_{n+1} z_r J^{up}_2\} | J^{up}_2 = k_2\}
\]

\[
= \lim_{m,n \to \infty} \prod_{s,u=0}^n \prod_{q,r=1}^m \sum_{k_1, k_2=0}^{\infty} e^{-\mu^q_i} \frac{\left(\mu^q_i\right)^k}{k_1!} e^{ia \zeta_{n+1} z_q k_1}
\]

\[
\cdot \sum_{k_2=0}^{\infty} e^{-\mu^{ur}_r} \frac{\left(\mu^{ur}_r\right)^k}{k_2!} e^{-ia \zeta_{n+1} z_r k_2}
\]

12 Recall: if \( \{X_n\}_{n \in \mathbb{N}^+} \) and \( X \) are random vectors in \( \mathbb{R}^k \), and \( \{\varphi_n\}_{n \in \mathbb{N}^+} \) and \( \varphi X \) are the associated characteristic functions, then (I) \( X_n \) converges in distribution to \( X \) if \( \varphi_{X_n}(s) \to \varphi_X(s) \) for all \( s \in \mathbb{R}^k \), and (II) if \( \varphi_X(s) \to \varphi(s) \) pointwise for all \( s \in \mathbb{R}^k \), and if \( \varphi \) is continuous on \( 0 \), then \( \varphi \) is the characteristic function of \( X \) and \( X_n \) converges in distribution to \( X \).

13 This discretization idea is inspired by Hanson (2007).
\[
\begin{align*}
&= \lim_{m,n \to \infty} \prod_{s=0}^{n} \prod_{q=1}^{m} \exp\{\mu_{1}^{\ast q}(e^{ia\zeta_{n+1}z_{q}} - 1)\} \exp\{\mu_{2}^{\ast r}(e^{-ia\zeta_{n+1}z_{r}} - 1)\} \\
&= \lim_{m,n \to \infty} \exp\left\{\sum_{s=0}^{n} \sum_{q=1}^{m} \mu_{1}^{\ast q}(e^{ia\zeta_{n+1}z_{q}} - 1)\right\} \exp\left\{\sum_{u=0}^{n} \sum_{r=1}^{m} \mu_{2}^{\ast r}(e^{-ia\zeta_{n+1}z_{r}} - 1)\right\}, \\
&= \exp\left\{\int_{0}^{t} (e^{ia\zeta_{t}z} - 1)\gamma_{s}F_{s}(dz)ds\right\} \exp\left\{\int_{0}^{t} (e^{-ia\zeta_{t}z} - 1)\gamma_{s}F_{s}(dz)ds\right\}, \\
&= \exp\left\{\int_{0}^{t} (\mathbb{E}[e^{ia\zeta_{t}Z_{s}}] + \mathbb{E}[e^{ia\zeta_{t}Z_{s}}^{\ast}] - 2)\gamma_{s}ds\right\},
\end{align*}
\]

where \(\ast\) designates the complex conjugate. Here, the fourth equality uses the law of total expectation, while the sixth equality uses the Taylor expansion of the exponential function. Finally, from the expansion

\[
\mathbb{E}[e^{ia\zeta_{t}Z_{s}}] = 1 + ia\zeta_{t}\mathbb{E}[Z_{s}^{1}] - \frac{1}{2}a^{2}\zeta_{t}^{2}\mathbb{E}[Z_{s}^{2}] - \frac{1}{6}ia^{3}\zeta_{t}^{3}\mathbb{E}[Z_{s}^{3}] + \mathcal{O}(a^{4}\zeta_{t}^{4}\mathbb{E}[Z_{s}^{4}]),
\]

we find after internal cancellation that

\[
\varphi_{L}(a) = \exp\left\{-a^{2}\zeta_{t}^{2}\int_{0}^{t} \mathbb{E}[Z_{s}^{2}]\gamma_{s}ds + \mathcal{O}\left(\frac{a^{4}\zeta_{t}^{4}\int_{0}^{t} \mathbb{E}[Z_{s}^{4}]\gamma_{s}ds}{1}\right)\right\}. 
\]

Upon specifying \(\zeta_{t}\) as in (A.1) and by noting that \(\int_{0}^{t} \gamma_{s}ds \leq \int_{0}^{t} \mathbb{E}[Z_{s}^{2}]\gamma_{s}ds\) (the jump sizes \(Z_{s}\) takes values in \(\mathbb{N}^{+}\)) and \(\int_{0}^{t} \mathbb{E}[Z_{s}^{4}]\gamma_{s}ds \leq \sup_{s \in [0,t]} \mathbb{E}[Z_{s}^{4}] \int_{0}^{t} \gamma_{s}ds\) (the fourth moment is bounded by assumption), letting \(\int_{0}^{t} \gamma_{s}ds \to \infty\) it follows that \(\Phi_{L}(a) \to \exp\left\{-\frac{1}{2}a^{2}\right\}\), which is the characteristic function of the standard normal distribution.

Proposition 1 may now immediately be established upon noting that for large \(\bar{\gamma}\) the approximation \(\zeta_{t}L_{t} \sim N(0,1)\). Hence, \(\sigma_{m}\zeta_{t}L_{t} \sim (\sigma_{m}/\sqrt{t})N(0,t)\). Rearranging this, the result follows.

**Appendix B. Proof of Proposition 3**

**Proof.** Following Barles and Souganidis (1991) and Fleming and Soner (2005), we first recall the following definitions, suitably adapted to the problem at hand

(1) (Monotonicity): For all \(\varepsilon \in \mathbb{R}_{+}\), and for all unit vectors \(\{\hat{e}_{i}\}_{i=1}^{2K+3}\) we require that

\[
\mathcal{G}[\theta^{n+1}] \leq \mathcal{G}[\theta^{n+1} + \varepsilon \hat{e}_{i}],
\]

where \([\hat{e}_{i}]_{j} = \delta_{ij}\) where \(\delta_{ij}\) is the Kronecker delta.
(2) (Stability): 
\[ \| \vartheta^n \|_\infty \leq C, \]  
for some constant \( C \) independent of \( \Delta t \), as \( \Delta t \to 0 \), where \( \| x \|_\infty \equiv \max_i |x_i| \).

(3) (Consistency): Finally, we require that the system satisfies the basic consistency condition 
\[ \lim_{\Delta t \to 0} \left| \frac{1}{\Delta t} \left( G[\vartheta^{n+1}] - \vartheta^T_{i,j} \right) - \mathcal{H}[\vartheta((n+1)\Delta t, i\sigma_m, q_{\min} + j)] \right| = 0, \]  
where \( \mathcal{H}[\cdot] \) is the HJB QVI, (13), centered at the coordinate 
\[ (t, s, q) = ((n+1)\Delta t, i\sigma_m, q_{\min} + j). \]

We will work through these items systematically, focusing on the case where the limit order equation applies (the market order equation trivially satisfies 1, 2, and 3):

(1) (Monotonicity): This is trivially true for all elements of \( \vartheta^{n+1} \) which only have a positive coefficient instantiation in \( G_{\Delta t}[\vartheta^{n+1}] \). In fact, the only nontrivial case is \( \theta^T_{i,j} \). Specifically, upon eliminating all superfluous terms, the monotonicity condition requires:
\[
0 \leq \varepsilon + \Delta t \left\{ -2\gamma_{n+1} \varepsilon + \sup_{\delta_{n+1} \in \mathbb{R}_+} \left[ -\varepsilon \lambda_{n+1} e^{-\kappa_{n+1} \delta_{n+1}} \right] 1_{\{ q_{\min} + j \geq \Delta_{n+1} \}} \right. \\
+ \sup_{\delta_{n+1} \in \mathbb{R}_+} \left[ -\varepsilon \lambda_{n+1} e^{-\kappa_{n+1} \delta_{n+1}} \right] 1_{\{ q_{\min} + j < \Delta_{n+1} \}} \left. \right\} \\
= \varepsilon - \Delta t \left\{ 2\gamma_{n+1} \varepsilon + \varepsilon \lambda_{n+1} \inf_{\delta_{n+1} \in \mathbb{R}_+} \left[ e^{-\kappa_{n+1} \delta_{n+1}^+} \right] 1_{\{ q_{\min} + j \geq \Delta_{n+1} \}} \right. \\
+ \varepsilon \lambda_{n+1} \inf_{\delta_{n+1} \in \mathbb{R}_+} \left[ e^{-\kappa_{n+1} \delta_{n+1}^-} \right] 1_{\{ q_{\min} + j < \Delta_{n+1} \}} \left. \right\}.
\]

Since 
\[ \inf_{\delta_{n+1}^+} e^{-\kappa_{n+1} \delta_{n+1}^+} = 0, \]
we are left with the constraint 
\[ 2\gamma_{n+1} \Delta t \leq 1, \quad \text{for } n = N - 1, N - 2, \ldots, 0. \]  
Indeed, numerical experiments corroborate that violating (B.4) leads to a failure in convergence of the algorithm. Note that the inverse
proportionality between $\gamma$ and $\Delta t$ in practice can lead to very fine grid spacings for realistic values of the jump rate intensity: hence, (B.4) significantly impedes the rapidity with which the algorithm can be executed.

(2) (Stability): Taking the $L^\infty$-norm of the HJB equation and using the triangle inequality we obtain

$$\|\vartheta_{i,j}^n\|_\infty \leq \|\vartheta_{i,j}^{n+1}\|_\infty + \Delta t\{\ldots\}\|_\infty,$$

where $\{\ldots\}$ signifies the content of the curly brackets in the definition of $\mathcal{G}_t$. Setting $\Delta t \to 0$ and using the fact that the terminal condition $\vartheta_{i,j}^N$ by assumption is bounded from above, we obtain the desired result.

(3) (Consistency): This follows immediately upon substituting in the definitions of $\mathcal{G}_t$ and $\mathcal{H}$ in Eq. (B.3) and using the Taylor expansion

$$\vartheta_{i,j}^n = \vartheta_{i,j}^{n+1} - \Delta t \partial_t \vartheta_{i,j}^{n+1} + \mathcal{O}(\Delta t^2).$$

To complete the proof, we deploy similar arguments for the grid points near the boundaries (recall the rescaling of probabilities such as to avoid the price process leaving the grid). Since this effectively is a recapitulation of what has already been established, we omit the details here. Finally, regarding the comparison principle, loosely put, what is required is that if $\theta_1$ and $\theta_2$ are two solutions of the HJB QVI with $\theta_1(T, s, q) \geq \theta_2(T, s, q)$, then we require $\theta_1(t, s, q) \geq \theta_2(t, s, q)$ for all $t$. For a more rigorous definition, see the viscosity references. \hfill \Box

References


Bachelier, 1900, Théorie de la Spéciation, Annales Scientiﬁques de l’É.N.S., 17, 21–86. Available at: http://archive.numdam.org/ARCHIVE/ASENS/ASENS_1900_3_17_/ASENS_1900_3_17_21_0/ASENS_1900_3_17_21_0.pdf.

Bacry, Mastromatteo and Muzy, 2015, Hawkes Processes in Finance, Market Microstructure and Liquidity 1(01), 1550005.


Runggaldier, W., 2003, Jump Diffusion Models, in *Handbook of Heavy Tailed Distributions in Finance*, Elsevier Science B.V.


